

cfbinout & xtdhazard: Control-Function Estimation of Binary-Outcome Models and the Discrete-Time Hazard Model

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Outline

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- ② Stata Implementation
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- ④ Simulations
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Motivation

- ▶ **Discrete-time hazard model** obvious approach to modeling transitions if covariates \mathbf{x}_{it} **change** values **frequently**
- ▶ E.g. Panel data (for instance annual surveys)
 - » New information about \mathbf{x}_{it} in each panel wave $t = 1, \dots, T$
 - » Information about whether transition has occurred ($\tau_i = t$) or has not yet occurred ($\tau_i > t$) at same frequency
- ▶ Models prob. of unit i of transitioning in period t conditionally on not yet having transitioned (hazard λ_{it})

$$\lambda_{it} = P(\tau_i = t | a_i, \mathbf{x}_{it}, \tau_i \geq t) = F(a_i + \mathbf{x}_{it}\boldsymbol{\beta})$$

- ▶ **Focus:**
 - » Transition into **absorbing state** (single-spell model)
 - » Model with **unobserved heterogeneity**/frailty a_i

Estimation of Discrete-Time Hazard Model

- ▶ In terms of estimation, **coincide with binary outcome models** such as logit, probit, and cloglog (Jenkins, 1995; Tutz and Schmid, 2016)
 - » Re-formulating model in terms of binary outcome y_{it}
 - » $y_{it} = 1$ if $\tau_i = t$ and $y_{it} = 0$ if $\tau_i > t$
 - » y_{it} with $t > \tau_i$ not considered (not informative about transition to absorbing state)
- ▶ Allows for modelling (unit-level **unobserved heterogeneity**/frailty, a_i) as **random effects** like in panel binary outcome models (xtlogit, xtprobit, xtcloglog)

Modelling Unobserved Heterogeneity

- ▶ Random effects do not accommodate correlation of unobserved heterogeneity and covariates
- ▶ **Linear probability model** (LPM), which allows for **fixed effects** accommodating such correlation, as possible alternative
- ▶ LPM with unit-level **fixed effects** heavily **biased** in single-spell hazard model setting (Farbmacher and Tauchmann, 2023)
 - » Not because of – possibly inappropriate – linear specification
 - » Not because of failure to eliminate unobserved heterogeneity (biased even in its absence)
- ▶ Cond. expectation of error term $\varepsilon_{it}^{\text{FE}}$ in fixed-effects model

$$E\left(\varepsilon_{it}^{\text{FE}} \mid a_i, \mathbf{x}_{i1}, \dots, \mathbf{x}_{iT}, \tau_i \geq t\right) = a_i + E(\bar{\mathbf{x}}_i)_t \boldsymbol{\beta}$$

Own-Differences IV Estimation

- ▶ Linear (internal) **instrumental variables** (IV) estimator to deal with unobserved heterogeneity (suggested by Farbmacher and Tauchmann, 2023)
 - » All explanatory variables \mathbf{x}_{it} instrumented by their own **(first) differences** $\Delta\mathbf{x}_{it}$
 - » Unlike fixed-effects in standard setting, does not accommodate arbitrary correlation between a_i and \mathbf{x}_{it}
 - » Rather assumption $\text{Cov}(\Delta\mathbf{x}_{it}, a_i) = \mathbf{0}$ required (while allowing for $\text{Cov}(\mathbf{x}_{it}, a_i) \neq \mathbf{0}$)
 - » Still subject to – some sort of – survivor bias (rather small in many settings)
 - » Does not suffer from bias caused by including unit-level fixed effects

Non-Linear Control Function Estimation

- ▶ Linear hazard specification (LPM) not too appealing
 - » At least if one thinks of LPM in terms of modelling a data generating process
- ▶ Rationale for using (first) own-differences as an instrument may also apply to nonlinear models
 - » Discussed in passing in Farbmacher and Tauchmann (2023)
 - » **Control function** approach (cf. Wooldridge, 2015)
 - » I.e. including first-stage residuals as additional second-stage regressors

Existing Stata Commands

- ▶ Instrumental variables and control function estimators (naturally) **implemented** in **Stata**
 1. `ivregress 2sls`
 2. `ivprobit`, `twostep`
 3. `ivreg210` (contributed by C.F. Baum, M.E. Schaffer, Steven Stillman)
 4. `ivcloglog` (contributed by W. Liu)
 - ... probably more
- ▶ Generating numerous internal instruments **cumbersome**
 - » In particular if factor variables syntax used
- ▶ Existing commands **not specific** to single spell hazard model setting
 - » User needs to check actively if data warrants using first-differences IV estimator
- ▶ No command for control function logit

Two New Stata Commands

1. **xtdhazard**

- » Checks data for being consistent with transition to absorbing state
- » Temporarily generates internal instruments
- » Calls either `ivregress 2sls` or `cfbinout`

2. **cfbinout**

- » Runs logit, probit, or cloglog control function estimator (following Wooldridge, 2015)
 - › Stata implementation follows Terza (2017)
- » If called by `xtdhazard` uses internal instruments (first- or higher-order differences)
- » Can also be used as stand-alone command using user-specified instruments

The `xtdhazard` Command

- ▶ Just a wrapper for more convenient IV estimation of discrete-time hazard model
- ▶ Straight-forward with estimator (two-part command name) `2sls` (calls `ivregress 2sls`)
 - » In terms of theory, in detail discussed in Farbmacher and Tauchmann (2023)
 - » Just `ivregress 2sls` in terms of Stata
- ▶ Less straight-forward with estimators `logit`, `probit`, or `cloglog` (calls `cfbinout`)
 - » Only some simulation-based discussion in online appendix to Farbmacher and Tauchmann (2023)
 - » Some issues that do not apply to linear model (binary rhs-variables, quasi-complete separation)

The `cfbinout` Command

- ▶ Implements control function estimators, i.e. first stage residuals included as additional regressors
- ▶ Follows and draws on Terza (2017) “Two-stage residual inclusion estimation: A practitioners guide to Stata implementation [st0505]”
 - » Considering ML estimation in second stage
- ▶ Allows for logit, probit, or cloglog link (two part command name `cfbinout link`)
 - » For probit link (largely) equivalent to `ivprobit`, `twostep`
- ▶ Estimates a first stage for all regressors if called by `xtdhazard`

Specific issues with `cfbinout`

- ▶ Discrete endogeneous variables
 - » Not accommodated by standard control function approach (e.g. not allowed with `ivprobit`, `twostep`)
 - » According to Wooldridge (2015) ‘average structural form’ can – under certain assumptions – still be estimated by including **generalized residuals** from **binary outcome first stage**
 - » `cfbinout` allows specifying the link-function for first stage (probit, logit, linear) and uses generalized residuals for the former
- ▶ Quasi complete separation
 - » First stage may be subject to quasi complete separation (prone to if `cfbinout` is called by `xtdhazard`)
 - » `cfbinout` optionally allowed to switch to linear first stage

Syntax of xtdhazard

```
xtdhazard estimator depvar indepvars [if] [in]  
[weight] [, options]
```

- ▶ Requires data to be xtset (declared panel data)
 - » *panelvar* and *timevar* required
- ▶ Estimators
 - (i) 2sls (→ linear 2SLS, calls `ivregress 2sls`)
 - (ii) logit (→ control function logit, calls `cfbinout`)
 - (iii) probit (→ control function probit, calls `cfbinout`)
 - (iv) cloglog (→ control function complementary log-log, calls `cfbinout`)

Selected Options for xtdhazard

► General Options

- » difference(*numlist*): Sets order of differencing
 - › `difference(1)`, i.e. (only) using first-differences as instruments, is the default
 - › Yet, for instance, also `difference(1 3)` is possible (yet makes probably little sense)
- » instruments(*varlist*): Specifies additional, non-internal instruments
- » noabsorbing: Forces estimation if `depvar` does not indicate absorbing state

► Options for estimator 2sls

- » interactinst: Use squares and interactions of instruments as additional instruments
- » nofirststage: Do not save first-stage coefficients in `e(G)` and do not perform checks regarding first-stage
- » underid(*string*): Calls `underid` from within `xtdhazard`

Selected Options for xtdhazard (cont.)

- ▶ Options for estimators logit, probit, and cloglog
 - » order(#): Specify order of control-function polynomial
 - > order(1), i.e. just including (generalized) residuals, is the default
 - > order(2), order(3), ... means also including higher powers (→ ivcloglog)
 - » noresgenerate: First-stage residuals are only temporarily generated; coefficients of first-stage residuals are not reported; first-stage coefficients are not saved
 - » resname(stub): First-stage residual permanently saved as *stub_varname2*; the default for *stub* is *res*
 - » replace: Variable *stub_varname2* replaced if already it exists

Syntax of cfbincout

```
cfbincout link depvar [varlist1]
(varlist2 = varlist_iv) [if] [in] [weight],
[options]
```

► Link functions

- (i) probit: normal CDF, $\Phi(\beta\mathbf{x})$
- (ii) logit: logistic CDF, $\frac{1}{1+\exp(-\beta\mathbf{x})}$
- (iii) cloglog: Gumbel CDF, $1 - \exp(-\exp(\beta\mathbf{x}))$

Selected Options for `cfbinout`

- ▶ Largely the same as for `xtdhazard logit` (and `xtdhazard probit/cloglog`)
- ▶ Options not available with `xtdhazard logit` (`xtdhazard probit/cloglog`)
 - » `fslink(name)`: Specifies first-stage link function (logit [default] probit, linear)
 - » `fsswitch`: Switch equation-wise to `fslink(linear)` in case of quasi-complete separation in first-stage

Data Generating Process for Outcome y_{it}

$$\begin{aligned}\lambda_{it} &= P(y_{it} = 1 | a_i, \mathbf{x}_{it}, \tau_i \geq t) \\ &= F(a_i + \beta_{con}x_{con,it} + \beta_{bin}x_{bin,it} + \beta_{tiv}x_{tiv,i})\end{aligned}$$

- ▶ $y_{it} = \text{"missing"}$ if $\tau_i < t$ (i.e. $y_{it-1} \neq 0$)
- ▶ a_i : unobserved heterogeneity (beta distr.)
- ▶ x_{con} : **continuous** (normal distr.)
- ▶ x_{bin} : **binary** (Bernoulli distr.)
- ▶ x_{tiv} : **time-invariant** and contin. (wt. sum of beta distr.)
- ▶ $F(\cdot) = \Phi(\cdot)$ (alternatively logistic CDF, Gumbel CDF)
- ▶ $\beta_{con}, \beta_{bin}, \beta_{tiv}$ take value of 1 (rescaled for logit and cloglog) or 0 (\rightarrow various **exclusion restrictions**)

Correlation of Unobserved Heterogeneity and \mathbf{x} Vars.

```
. correlate a x_tiv x_bin x_con d.x_bin d.x_con
(obs=16,000)
```

	a	x_tiv	x_bin	x_con	D. x_bin	D. x_con
a	1.0000					
x_tiv	-0.7517	1.0000				
x_bin	-0.5843	0.4349	1.0000			
x_con	-0.8131	0.6103	0.4727	1.0000		
x_bin					1.0000	
D1.	0.0034	-0.0000	0.5698	-0.0053	1.0000	
x_con						1.0000
D1.	-0.0028	0.0018	-0.0054	0.4134	-0.0103	1.0000

- ▶ **Unobserved heterogeneity a** (negatively) **correlated** with all \mathbf{x} variables, yet **uncorrelated** with $\Delta\mathbf{x}$
- ▶ \mathbf{x} variables **stationary** (time series properties matter Farbmacher and Tauchmann, 2023)

Coefficient Values and Sample

- ▶ $N = 4000$ (# of units), $T = 5$ (# of periods)
- ▶ Patters of included \mathbf{x} vars.
 - (i) Only continuous x_{con}
 - (ii) Continuous x_{con} and time-invariant x_{tiv}
 - (iii) Continuous x_{con} and binary x_{bin}
 - (iv) All three x_{con} , x_{tiv} , x_{bin}
- ▶ Exclusion restrictions taken into account in estimation
- ▶ Average hazard $\bar{\lambda} = 0.25$ (alternatively $\bar{\lambda} = 0.05$)

▶ distribution of λ

Estimators and Replications

- ▶ Three estimating procedures
 - (i) 'Naive' probit (alternatively logit and cloglog)
 - (ii) Linear 2SLS (\rightarrow `xtdhazard 2sls`)
 - (iii) Control function probit (\rightarrow `xtdhazard probit`, alternatively `xtdhazard logit` and `xtdhazard cloglog`)
- ▶ Focus on (sample) average partial effects of x_{con} and x_{bin}
 - » β_{tiv} not identified in 2SLS and control function estimation
- ▶ Monte Carlo simulations using 2000 replications

Simulation Results: Probit, $\bar{\lambda} = 0.25$

Table: Average Partial Effects (probit, $\bar{\lambda} = 0.25$)

Av. partial effect of	Estimator	Scenarios: rhs vars. inclusion			
		X_{con}	X_{con}, X_{tiv}	X_{con}, X_{bin}	$X_{con}, X_{bin}, X_{tiv}$
X_{con}	Naive probit	0.106	0.153	0.113	0.126
	CF probit	0.306	0.282	0.249	0.214
	2SLS	0.306	0.282	0.249	0.216
	True value	0.305	0.282	0.260	0.219
X_{bin}	Naive probit			0.219	0.205
	CF probit			0.282	0.248
	2SLS			0.280	0.250
	True value			0.284	0.248

Note: The table presents simulation results comparing estimated and true mean-marginal effects in a discrete time hazard framework. The estimations were based on 2000 simulation runs. "True value" refers to the actual mean-marginal effects used in the simulation. The constants chosen for different scenarios were selected to achieve a mean probability of 0.25.

Simulation Results: Alternative Simulations

Pattern of results generally the same:

- (i) logit estimation, $\bar{\lambda} = 0.25$ ▶ logit, $\bar{\lambda} = 0.25$
- (ii) cloglog estimation, $\bar{\lambda} = 0.25$ ▶ cloglog, $\bar{\lambda} = 0.25$
- (iii) probit estimation, $\bar{\lambda} = 0.05$ ▶ probit, $\bar{\lambda} = 0.05$

Real Data Application: Replication of Cantoni (2012)

Cantoni (2012, EJ): “Adopting a New Religion: The Case of Protestantism in 16th Century Germany”

- ▶ Research Question: Which factors explain territories adopting protestantism?
- ▶ Uses historical panel data (74 territories in 16th century Germany; 5 years from 1532 to 1600)
- ▶ Key results: (i) distance to Wittenberg and (ii) neighbours' religious choices matter for adoption of protestantism
- ▶ Focus on fixed-effects specification Cantoni (2012, p. 522, Table 6, Column 3)
 - » Subject to minor change regarding clustering

Original Result by Cantoni (2012)

```
. xtreg refuntil lagrefneighbors, cluster(kreis) nonest fe
```

Fixed-effects (within) regression

Number of obs = 370

Group variable: tcode

Number of groups = 74

R-squared:

Within = 0.2348

Between = 0.0751

Overall = 0.0974

Obs per group:

min = 5

avg = 5.0

max = 5

F(1, 9) = 37.76

corr(u_i, Xb) = 0.0471

Prob > F = 0.0002

(Std. err. adjusted for 10 clusters in kreis)

refuntil	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lagrefneig_s	.6792987	.1105446	6.15	0.000	.4292295	.9293678
_cons	.1832339	.033076	5.54	0.000	.1084108	.2580569
sigma_u	.4024456					
sigma_e	.26168501					
rho	.70283546	(fraction of variance due to u_i)				

- ▶ Neighbouring territories' confession (lagrefneighbors) matters

How is Absorbing State Dealt With?

```
. gen l_refuntil = l.refuntil
(74 missing values generated)

. tab refuntil l_refuntil if e(sample)
```

refuntil	l_refuntil		Total
	0	1	
0	166	0	166
1	28	102	130
Total	194	102	296

- ▶ “Selection into Protestantism was effectively an absorbing state” (Cantoni, 2012, p. 523)
- ▶ Not taken into account in Fixed Effects Specification

Territories at Risk Only

```
. xtreg refuntil lagrefneighbors if L_refuntil != 1, cluster(kreis) fe
```

```
Fixed-effects (within) regression      Number of obs   =      268
Group variable: tcode                  Number of groups =       74

R-squared:                             Obs per group:
    Within = 0.1442                      min =          1
    Between = 0.8669                     avg =          3.6
    Overall = 0.0004                      max =          5

corr(u_i, Xb) = -0.4007                  F(1, 9)         =      10.27
                                          Prob > F        =      0.0107
```

(Std. err. adjusted for 10 clusters in kreis)

```
-----+-----
      refuntil |          Coef.   Robust
              |          Std. Err.   t   P>|t|   [95% Conf. Interval]
-----+-----
lagrefneig_s |   .5722313   .1785385   3.21  0.011   .1683491   .9761135
   _cons     |   .0104711   .044465   0.24  0.819   -.0901157   .1110578
-----+-----
sigma_u      |   .43093158
sigma_e      |   .28265007
rho          |   .6991975   (fraction of variance due to u_i)
-----+-----
```

► Only considering territories at risk makes little difference

Linear First-Differences IV Estimation

```
. xtdhazard 2sls refuntil lagrefneighbors if l_refuntil != 1, cluster(kreis)
> underid(underid)
```

```
Linear discrete-time hazard model      Number of obs      =      194
first-differences IV estimation        Number of groups   =       61
                                       Wald chi2(1)       =       0.25
                                       Prob > chi2        =      0.619
                                       R-sq              =       .
```

(Std. err. adjusted for 10 clusters in kreis)

```
-----+-----
      refuntil |          Coef.   Clustered
              |          Std. Err.   z   P>|z|   [95% Conf. Interval]
-----+-----
lagrefneig_s |  -.7880119   1.582831   -0.50   0.619   -3.890303   2.31428
   _cons     |   .4154438   .5556795    0.75   0.455   -.6736681   1.504556
-----+-----
```

```
Underidentification test: j =      6.35; Chi-sq( 1); p-value = 0.0117
```

- ▶ xtdhazard 2sls used for estimation
- ▶ Does not confirm that neighbours' religious choices matter
- ▶ Underidentification (d.lagrefneig_s weak instr.) rejected

Control Function Logit Estimation

```
. xtdhazard logit refuntil lagrefneighbors if l_refuntil != 1, cluster(kreis)
```

```
Logit discrete-time hazard model      Number of obs      =      194
first-differences CF estimation        Number of groups    =       61
                                        Wald chi2(1)        =       0.10
                                        Prob > chi2         =       0.749
```

```
Log pseudolikelihood = -79.136
```

(Std. err. adjusted for 10 clusters in kreis)

refuntil	Coef.	Clustered Std. Err.	z	P> z	[95% Conf. Interval]	
lagrefneig_s	-5.736172	17.90578	-0.32	0.749	-40.83086	29.35851
res_lagref_s	6.672166	15.71915	0.42	0.671	-24.1368	37.48113
_cons	.1694307	6.219791	0.03	0.978	-12.02113	12.36

- ▶ xtdhazard logit used for estimation
- ▶ Results qualitatively the same as counterparts from 2SLS

Control Function Logit Estimation (cont.)

```
. margins, dydx(lagrefneighbors)
```

```
Average marginal effects
Model VCE: Clustered
```

```
Number of obs = 194
```

```
Expression: Pr(refuntil), predict()
dy/dx wrt: lagrefneighbors
```

		Delta-method				[95% Conf. Interval]	
	dy/dx	Std. Err.	z	P> z			
lagrefneig_s	-.7007748	2.234062	-0.31	0.754	-5.079456	3.677906	

- ▶ Result in terms of average marginal effects

Control Function Cloglog Estimation

```
. xtdhazard cloglog refuntil lagrefneighbors if l_refuntil != 1, cluster(kreis)
> difference(2) replace
```

```
Cloglog discrete-time hazard model      Number of obs      =      133
2nd-differences CF estimation           Number of groups   =       50
                                         Wald chi2(1)       =       0.53
                                         Prob > chi2        =       0.467

Log pseudolikelihood =   -50.153
```

(Std. err. adjusted for 10 clusters in kreis)

refuntil	Coef.	Clustered Std. Err.	z	P> z	[95% Conf. Interval]	
lagrefneig_s	-4.825455	6.626972	-0.73	0.467	-17.81408	8.163171
res_lagref_s	7.00504	9.303846	0.75	0.451	-11.23016	25.24024
_cons	-.0397897	2.861651	-0.01	0.989	-5.648523	5.568944

- ▶ xtdhazard cloglog used for estimation
- ▶ Difference-in-Differences used as instrument

Control Function Cloglog Estimation (cont.)

```
. margins, dydx(lagrefneighbors)
```

```
Average marginal effects
Model VCE: Clustered
```

```
Number of obs = 133
```

```
Expression: Pr(refuntil), predict()
dy/dx wrt: lagrefneighbors
```

	Delta-method					
	dy/dx	Std. Err.	z	P> z	[95% Conf. Interval]	
lagrefneig_s	-.5720407	.8870068	-0.64	0.519	-2.310542	1.166461

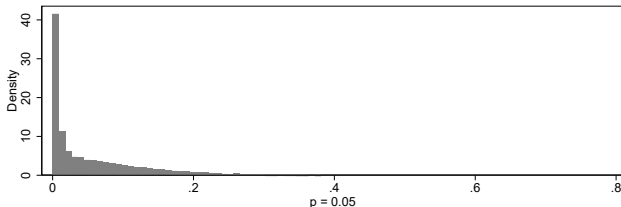
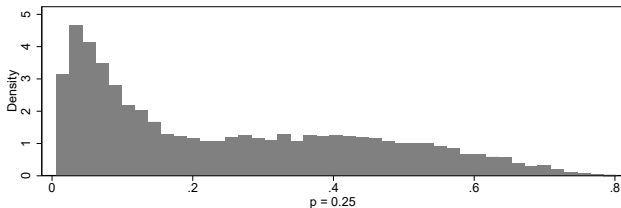
- ▶ Result in terms of average marginal effect
- ▶ Point estimate changes only marginally

Conclusions

- ▶ Using 2SLS or non-linear control function estimation probably better suited for dealing with unobserved heterogeneity in a single-spell hazard setting (than including unit fixed effects)
- ▶ These well-known estimators already implemented in Stata
- ▶ Using numerous internal instruments renders implementation through existing commands cumbersome
- ▶ `xtdhazard` eases using these estimators for Stata users
- ▶ `cfbinout` can be used as stand-alone estimator, which complements `ivprobit` and `ivcloglog`

Distribution of Hazard

- ▶ Comparison of simulations with mean $\bar{\lambda} = 0.25$ and $\bar{\lambda} = 0.05$



Simulation Results: Logit, $\bar{\lambda} = 0.25$

Table: Average Partial Effects (logit, $\bar{\lambda} = 0.25$)

Av. partial effect of		Scenarios: rhs vars. inclusion			
		Estimator	X_{con}	X_{con}, X_{tiv}	X_{con}, X_{bin}
X_{con}	Naive logit	0.035	0.055	0.044	0.058
	CF logit	0.102	0.101	0.099	0.097
	2SLS	0.102	0.101	0.099	0.097
	True value	0.103	0.102	0.101	0.098
X_{bin}	Naive logit			0.079	0.083
	CF logit			0.102	0.100
	2SLS			0.102	0.100
	True value			0.102	0.100

Note: The table presents simulation results comparing estimated and true mean-marginal effects in a discrete time hazard framework. The estimations were based on 2000 simulation runs. "True value" refers to the actual mean-marginal effects used in the simulation. The constants chosen for different scenarios were selected to achieve a mean probability of 0.25.

Simulation Results: Cloglog, $\bar{\lambda} = 0.25$

Table: Average Partial Effects (cloglog, $\bar{\lambda} = 0.25$)

Av. partial effect of	Estimator	Scenarios: rhs vars. inclusion			
		X_{con}	X_{con}, X_{tiv}	X_{con}, X_{bin}	$X_{con}, X_{bin}, X_{tiv}$
X_{con}	Naive cloglog	0.057	0.088	0.070	0.088
	CF cloglog	0.166	0.161	0.156	0.148
	2SLS	0.166	0.161	0.156	0.148
	True value	0.167	0.163	0.160	0.150
X_{bin}	Naive cloglog			0.125	0.127
	CF cloglog			0.162	0.153
	2SLS			0.162	0.153
	True value			0.162	0.153

Note: The table presents simulation results comparing estimated and true mean-marginal effects in a discrete time hazard framework. The estimations were based on 2000 simulation runs. "True value" refers to the actual mean-marginal effects used in the simulation. The constants chosen for different scenarios were selected to achieve a mean probability of 0.25.

Simulation Results: Probit, $\bar{\lambda} = 0.05$

Table: Average Partial Effects (probit, $\bar{\lambda} = 0.05$)

Av. partial effect of	Estimator	Scenarios: rhs vars. inclusion			
		X_{con}	X_{con}, X_{tiv}	X_{con}, X_{bin}	$X_{con}, X_{bin}, X_{tiv}$
X_{con}	Naive probit	0.035	0.052	0.039	0.050
	CF probit	0.102	0.095	0.087	0.083
	2SLS	0.101	0.094	0.086	0.083
	True value	0.104	0.096	0.089	0.084
X_{bin}	Naive probit			0.060	0.057
	CF probit			0.080	0.069
	2SLS			0.080	0.069
	True value			0.080	0.069

Note: The table presents simulation results comparing estimated and true mean-marginal effects in a discrete time hazard framework. The estimations were based on 2000 simulation runs. "True value" refers to the actual mean-marginal effects used in the simulation. The constants chosen for different scenarios were selected to achieve a mean probability of 0.05.

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